

# Bayes spaces: use of improper priors and distances between densities

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# Distances/divergences for probability densities

## applications:

- goodness of fit
- fitting distributions (e.g. kernel estimation)
- information theory (e.g. Kullback-Leibler div.)

## examples of distances/divergences:

- from functional spaces:  $L^1$ ,  $L^2$ ,  $L^\infty$  applied to pdf's or cdf's
- ad-hoc: Hellinger-Matusita; Chi-square;
- from information theory: Kullback-Leibler, ...

# What is lacking in these distances/divergences?

## compatibility with probabilistic operations

### two relevant operations:

- **convolution of pdf's**: associated with sum of random variables.
- **Bayes updating**: information acquisition

there is a need of a meaningful algebraic/geometric structure associated with Bayes updating

# simplex and compositional data

**composition:** equivalent class of real vectors with proportional positive components

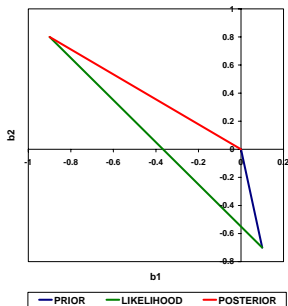
- components quantify parts of a whole
- only ratios between components are informative
- standard representative: a point in the simplex (components adding to 1)

## Euclidean structure of the simplex

- interpretable operations: perturbation  $\oplus$ , powering  $\odot$
- Aitchison metrics: inner product, norm and distance
- orthogonal bases, reference measure

**perturbation in the simplex is the Bayes formula for discrete probability vectors**

# Coordinate representation

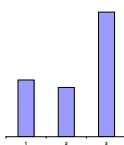


simplex coordinates

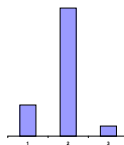
balance-coordinates

$$b_1 = \frac{1}{\sqrt{2}} \log \frac{p_1}{p_2}$$

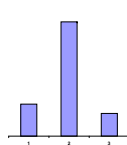
$$b_2 = \sqrt{\frac{2}{3}} \log \frac{(p_1 p_2)^{1/2}}{p_3}$$



prior



likelihood



posterior

# leading ideas and goal

**heuristic idea:** a histogram is a composition and it is equivalent to a simplex element

- increasing the number of classes in a histogram it approaches a pdf...
- perturbation in the simplex is discrete Bayes formula, it can be extended to the continuous case...

## goal

- vector space structure with Bayes updating as addition
- metric spaces of densities
- Hilbert spaces of densities

# $\lambda$ -equivalent measures

## assuring existence of densities

measurable space:  $(\Omega, \mathfrak{A})$

sigma-additive measures:  $\lambda$  equivalent to  $\mu$  (finite or infinite)

$$\mu \equiv \lambda \Leftrightarrow \forall A \in \mathfrak{A}, (\lambda(A) = 0 \Leftrightarrow \mu(A) = 0)$$

- $\lambda$  and  $\mu$  have the same support
- the Radon-Nikodým derivative (density) exists

$$\frac{d\mu}{d\lambda} = f_\mu$$

## Examples of equivalent measures:

- $\Omega = \mathbb{R}$ : normal, t-student, Lebesgue measure (improper uniform)
- $\Omega = \mathbb{R}_+$ : log-normal, gamma, Lebesgue measure in  $\mathbb{R}_+$
- $\Omega = \{0, 1, 2, 3, \dots\}$ : Poisson, geometric, counting measure

# B-equivalence: proportional densities

**reference measure**  $\lambda$ :  $\mu_1, \mu_2 \equiv \lambda$

**densities:**  $f_1 = d\mu_1/d\lambda, f_2 = d\mu_2/d\lambda$

$\mu_1, \mu_2$  are  $B$ -equivalent,  $\mu_1 =_B \mu_2$ , iff

$$\exists c > 0, \forall A \in \mathfrak{A}, \mu_1(A) = c \cdot \mu_2(A), f_1 = c \cdot f_2, (\lambda - a.e.)$$

## Remarks

- **likelihood principle:** proportional likelihood functions convey identical information
- **normalization of probabilities:** not essential
- **essential information:** ratios of probabilities



# perturbation and powering

**Bayes space, reference  $\lambda$ :** elements of  $B(\lambda)$  are classes of  $B$ -equivalent measures/densities

**perturbation (addition, group operation):**  $f_1, f_2 \in B(\lambda)$

$$f_1 \oplus f_2 =_B f_1 \cdot f_2 \ (\lambda - a.e.) \quad , \quad (\mu_1 \oplus \mu_2)(A) = \int_A \frac{d\mu_1}{d\lambda} \frac{d\mu_2}{d\lambda} d\lambda$$

**powering (multiplication):**  $f \in B(\lambda), \alpha \in \mathbb{R}$

$$\alpha \odot f =_B f^\alpha \ (\lambda - a.e.) \quad , \quad (\alpha \odot \mu)(A) = \int_A \left( \frac{d\mu}{d\lambda} \right)^\alpha d\lambda$$

# vector space and Bayes theorem

## $B(\lambda)$ -space includes

- **prior densities**, proper or improper
- **likelihood functions**, integrable or not
- **posterior densities**, proper or improper

## Bayes theorem

$$\rho =_B L \oplus \pi =_B \left( \bigoplus_{i=1}^n L_i \right) \oplus \pi$$

**perturbation/Bayes updating** is an internal operation in  $B(\lambda)$

**powering** means (linear) weighting

**iterate perturbation:** group properties allow improper intermediate steps

# vector space and exponential families

**exponential families**,  $k$ -parametric, natural parameters

$$f(x|\vec{\theta}) = C(\vec{\theta})g(x) \exp \left[ \sum_{j=1}^k \theta_j T_j(x) \right]$$

$B(\lambda)$  **expression**

$$f_{\lambda}(x) =_B g_{\lambda}(x) \oplus \bigoplus_{j=1}^k (\theta_j \odot \exp[T_j(x)])$$

**$k$ -dimensional affine subspace**

- $g_{\lambda}(x)$  origin of the affine subspace
- $\exp[T_j(x)]$  basis of the subspace
- $\theta_j$  coordinates of  $f_{\lambda}(x)$

**probability densities** are a convex cone of the  $k$ -dimensional affine subspace

# example: distribution of a sample maximum

$n$ -sample, distribution  $F$ , density  $f$   
**density of sample maximum**

$$f_M(x) = nf(x) \cdot [F(x)]^{n-1}$$

reference  $f_0$

$$f_M(x) =_B \underbrace{\frac{f(x)}{f_0(x)}}_{\text{origin}} \oplus (n-1) \odot \underbrace{[F(x)]}_{\text{direction}}$$

**the family is 1-parametric and follows a straight-line with  $n$**

# centered log-ratio mapping

## clr for compositions

$$\text{clr}(\vec{x}) = \log(x_1, x_2, \dots, x_k) - \frac{1}{k} \sum_{j=1}^k \log x_j \quad , \quad \sum_{j=1}^k \text{clr}_j(\vec{x}) = 0$$

**clr in  $B(P)$  reference**  $P$  probability measure; density  $f_P$   
**definition of clr**,  $f \in B(P)$

$$\text{clr}(f) = \log(f) - \frac{1}{P(\Omega)} \int \log(f(x)) f_P(x) dx$$

$P$  prob. measure  $\Leftrightarrow P(\Omega) = 1$

**clr mapping is linear; scale and  $B(P)$ -reference invariant**

# $B^q(P)$ spaces

$B^q(P)$  space of measures/densities,  $1 \leq q < \infty$

$$B^q(P) = \left\{ f \in B(P) : \int |\log f(x)|^q f_P(x) dx < +\infty \right\},$$

- clr exists for densities in  $B^1(P)$ ;
- $clr : B^1(P) \rightarrow L_0^1(P)$  is one-to-one
- $B^1(P) \supseteq B^2(P) \supseteq \dots \supseteq B^\infty(P)$
- $B^q(P)$  are Minkowsky metric spaces

$$d_{B^q}(f_1, f_2) = d_{L^q}(clr(f_1), clr(f_2)) = \left[ \int (clr(f_1) - clr(f_2))^q dP \right]^{1/q}$$

# B-derivative

$f(x|t) \in B^1(P)$ ;  $t$  external variable (time, space, sample values)

$f : \mathbb{R} \rightarrow B^1(P)$

## Definition of B-derivative

$$\frac{d^\oplus}{dt} f(x|t) =_B \lim_{h \rightarrow 0} \frac{1}{h} \odot [f(x|t+h) \ominus f(x|t)]$$

if it exists.  $\ominus = \oplus(-1) \odot$

- describes change of densities with  $t$
- differential calculus and differential equations for densities/measures
- useful concept in applications (Bayesian, robust stats.)

# Hilbert space

$B^2(P)$  is a separable Hilbert space

$$\text{clr} : B^2(P) \leftrightarrow L_0^2(P)$$

**inner product**,  $f_1, f_2 \in B^2(P)$

$$\langle f_1, f_2 \rangle_{B^2} = \langle \text{clr}(f_1), \text{clr}(f_2) \rangle_{L^2}$$

**distance and norm**

$$d_{B^2}(f_1, f_2) = d_{L^2}(\text{clr}(f_1), \text{clr}(f_2)) \quad , \quad \|f_1\|_{B^2} = \|\text{clr}(f_1)\|_{L^2}$$



# Hilbert basis and Fourier coordinates

$\psi_0, \psi_1, \psi_2, \dots$  a Hilbert basis in  $L^2(P)$

$\psi_0(x)$  constant function

**Hilbert basis of  $B^2(P)$**

$\exp(\psi_1), \exp(\psi_2), \dots$

**coordinates: Fourier coefficients**,  $f \in B^2(P)$

$$f =_B \bigoplus_{j=1}^{\infty} \langle f, \exp(\psi_j) \rangle_{B^2} \odot \exp(\psi_j)$$

- Fourier coefficients are real orthogonal coordinates
- if normalized, distances, norms, orthogonal projections, ... are computed as  $\ell^2$  sequences
- they allow to use standard "real multivariate statistics"

# the Normal family

**reference**  $P = N(0, 1)$  with Lebesgue density  $f_0$   
 **$P$ -density**  $g$  corresponding to  $N(m, \sigma^2)$

$$g(x) =_B f(x)/f_0(x) =_B \exp\left(-\frac{(x-m)^2 - \sigma^2 x^2}{2\sigma^2}\right)$$

**clr**

$$\text{clr}(g)(x) = \frac{x^2 - 1}{2} + \frac{1 - x^2 + 2mx}{2\sigma^2}$$

**distance**,  $g_i \sim N(m_i, \sigma_i^2)$

$$d_{B^2}^2(g_1, g_2) = \frac{1}{2} \left( \frac{1}{\sigma_1^2} - \frac{1}{\sigma_2^2} \right)^2 + \left( \frac{m_1}{\sigma_1^2} - \frac{m_2}{\sigma_2^2} \right)^2$$

# Fourier expansion of Normal family

**orthonormal Hilbert basis in  $L^2(N(0, 1))$ : Hermite**

$$\frac{1}{\sqrt{2\pi}} \int_{\mathbb{R}} H_j(2^{-1/2}x) H_k(2^{-1/2}x) e^{-x^2/2} dx = \delta_{jk} K_j^{-2}, \quad K_j = 2^{-j/2} (j!)^{-1/2}$$

**Hilbert basis in  $B^2(N(0, 1))$**

$$\exp[\psi_j(x)] = \exp[K_j H_j(2^{-1/2}x)], \quad j = 1, 2, \dots$$

**Fourier expansion**

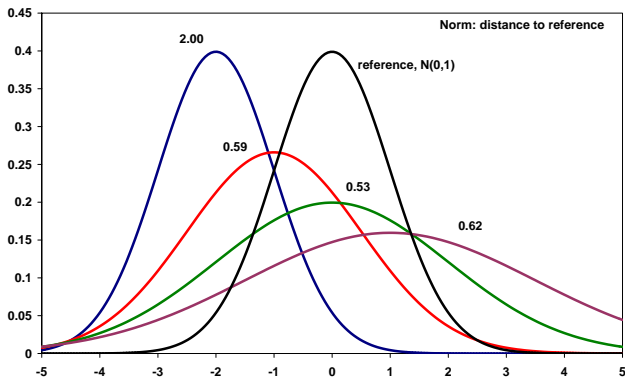
$$g(x) =_B c_1 \odot \exp[\psi_1(x)] \oplus c_2 \odot \exp[\psi_2(x)]$$

$$c_1 = \frac{m}{\sigma^2}, \quad c_2 = -\sqrt{2} \left( \frac{1}{2\sigma^2} - \frac{1}{2} \right), \quad c_j = 0, \quad j = 3, 4, \dots,$$

# norms of normals

reference:  $N(0, 1)$

$N(-2, 1)$ ,  $N(-1, 1.5^2)$ ,  $N(0, 2^2)$ ,  $N(1, 2.5^2)$



# conclusions

## results

- **proportional densities are considered equivalent ( $B(\lambda)$ )**
  - perturbation is the (extended) Bayes updating
  - proper and improper priors, likelihoods and posteriors are in  $B(\lambda)$
  - $(B(\lambda), \oplus, \odot)$  is a vector space
  - linear affine subspaces contain exponential families
- **$q$ -log-integrable densities ( $B^q(P)$ ) are metric spaces**
  - $clr : B^1(P) \rightarrow L_0^1(P)$  is one-to-one (isometry)
- **2-log-integrable densities in  $B^2(P)$  are a separable Hilbert space. Standard tools are then available:**
  - Hilbert basis and Fourier expansions
  - distances, norm, orthogonal projections
  - Aitchison geometry of the simplex is a particular case.

# conclusions

## consequences

- the new framework allows to rephrase most standard probabilistic models (Bayes theorem, exponential families, ...) in a simple and formal way
- tools of vector, metric and Hilbert spaces are now available for probabilistic/statistical modelling

## a good deal of research is still pending...

- the role of references in  $B^2(P)$
- possible uses of Fourier transforms in  $B^1(P)$  and  $B^2(P)$
- asymptotic theory on  $B(\lambda)$
- characteristics of well-known families (normal, gamma, beta, t-student,...)
- approximation of  $B^2(P)$  spaces by the simplex geometry

## some steps towards B-spaces

[Aitchison, J. \(1986\)](#). The Statistical Analysis of Compositional Data. Monographs on Statistics and Applied Probability. Chapman & Hall Ltd., London (UK). 416 p.

[Pawlowsky-Glahn, V. and J.J. Egozcue \(2001\)](#) Geometric Approach to Statistical Analysis on the Simplex, Stochastic Environmental Research and Risk Assessment 15 (5), 384–398

[Egozcue, J. J., J. L. Díaz-Barrero and V. Pawlowsky-Glahn \(2006\)](#). Hilbert Space of Probability Density Functions Based on Aitchison Geometry, Acta Mathematica Sinica 22 (4), 1175–1182

[van den Boogaart, K. G., J. J. Egozcue, and V. Pawlowsky-Glahn \(2010\)](#). Bayes linear spaces. Statistics and Operations Research Transactions, SORT 34 (2), 201–222

# example: zero-inflated Poisson exponential family

**reference measure:** counting measure

$$\nu(x) = 1, x = 0, 1, 2, \dots$$

**mixture expression**

$$f(x|\phi, p) = (1 - p) \cdot \delta(x) + p \cdot \frac{\phi^x e^{-\phi}}{x!}$$

$B(\nu)$  **2-parametric exponential family**

$$f(x|\theta_1, \theta_2) =_{B(\nu)} \underbrace{\frac{1}{x!}}_{\text{origin}} \oplus \left( \theta_1 \odot \underbrace{e^x}_{\text{basis}_1} \right) \oplus \left( \theta_2 \odot \underbrace{e^{\delta(x)}}_{\text{basis}_2} \right)$$

$$\theta_1 = \log \phi, \theta_2 = \log [(1 - p)e^\phi + p], C(\theta_1, \theta_2) = [\exp(\theta_2) + \exp(\exp(\theta_1)) - 1]^{-1}$$

**3-parametric conjugate family**

$$\pi(\theta_1, \theta_2) =_B \exp(t_0 \log C(\theta_1, \theta_2) + t_1 \theta_1 + t_2 \theta_2)$$